Sample Size Considerations for Comparing Dynamic Treatment Regimens in a SMART with a Longitudinal Outcome

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Joint with K.M. Kidwell, J.R. McKay, I. Nahum-Shani, T. Wu, D. Almirall

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[.] McKay, J. R., et al. (2015). J. Consult. Clin. Psychol.

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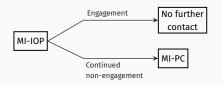
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This is a question about a sequence of treatments.

[.] McKay, J. R., et al. (2015). J. Consult. Clin. Psychol.

Dynamic treatment regimens (DTRs) operationalize clinical decision-making by recommending particular treatments to certain subsets of patients at specific times.



- MI-IOP: 2 motivational interviews to re-engage patient in intensive outpatient program
- **MI-PC**: 2 motivational interviews to engage patient in treatment of their choice.

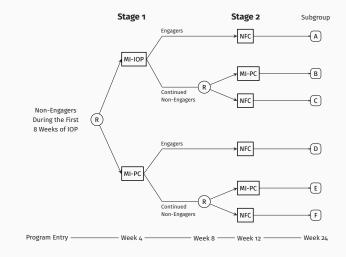
[.] Chakraborty, B., and E. E. M. Moodie (2013). Statistical Methods for Dynamic Treatment Regimes.

A **SMART** is one type of randomized trial design that can be used to answer questions at multiple stages of the development of a high-quality DTR.

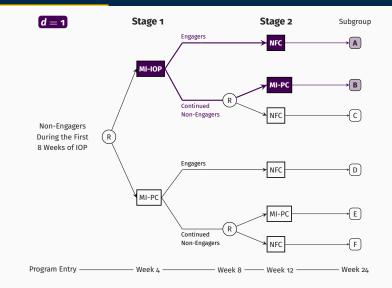
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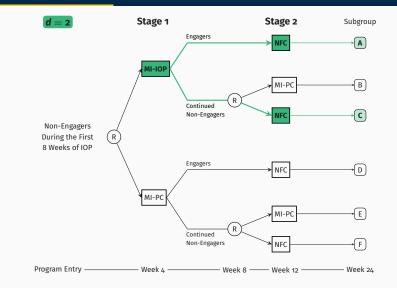
The key feature of a SMART is that some (or all) participants are randomized *more than once*.

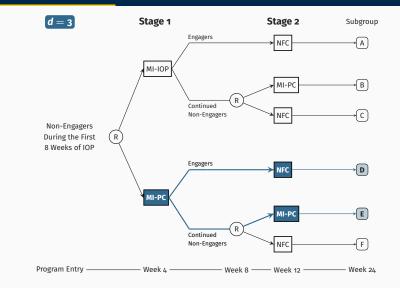
Motivating Example: The ENGAGE Study

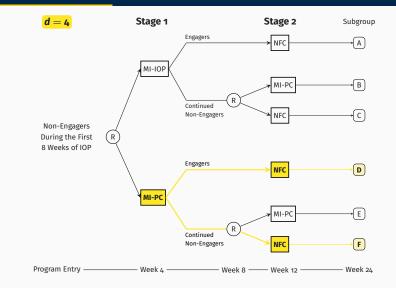


. McKay, J. R., et al. (2015). J. Consult. Clin. Psychol.



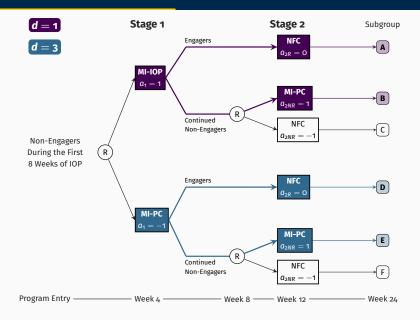




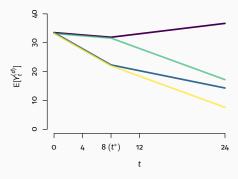


A common primary aim in a SMART is the comparison of two embedded DTRs using a continuous outcome collected at the end of the study.

Primary Aim



An Example Model for a Continuous Longitudinal Outcome in ENGAGE (Lu et al. 2016)



	<i>d</i> = 1	d = 2	d = 3	d = 4
a 1	1	1	-1	-1
a _{2R}	0	0	0	0
a _{2NR}	1	-1	1	-1

$$\begin{split} \mathsf{E} \left[\mathsf{Y}_{t}^{(d)} \mid \mathbf{X} \right] &:= \mu^{(d)}(\mathbf{X}_{i}; \eta, \gamma) \\ &= \eta^{\top} \mathbf{X}_{i} + \gamma_{\mathrm{o}} \\ &+ \mathbb{1} \{ t \leq t^{*} \} \left\{ \gamma_{1} t + \gamma_{2} a_{1} t \right\} \\ &+ \mathbb{1} \{ t > t^{*} \} \left\{ t^{*} \gamma_{1} + t^{*} \gamma_{2} a_{1} \\ &+ \gamma_{3} (t - t^{*}) + \gamma_{4} (t - t^{*}) a_{1} \\ &+ \gamma_{5} (t - t^{*}) a_{2NR} \\ &+ \gamma_{6} (t - t^{*}) a_{1} a_{2NR} \} \end{split}$$

8

"GEE-Type" Estimating Equations for Model Parameters

$$0 = \sum_{i=1}^{N} \sum_{d} \left[\underbrace{\frac{I^{(d)}(A_{1,i}, R_{i}, A_{2,i})}{P(A_{1,i} = a_{1})P(A_{2,i} = a_{2} \mid A_{1,i} = a_{1}, R_{i})}_{W^{(d)}(A_{1,i}, R_{i}, A_{2,i})} \cdot \mathbf{D}^{(d)}(\mathbf{X}_{i})^{\top} \cdot \mathbf{V}^{(d)}(\tau)^{-1} \cdot \left(\mathbf{Y}_{i} - \boldsymbol{\mu}^{(d)}(\mathbf{X}_{i}; \eta, \gamma)\right) \right],$$

where

- d specifies an embedded DTR,
- $W^{(d)}(A_{1,i}, R_i, A_{2,i}) = \mathbb{1}\{A_{1,i} = a_1\} (2R_i + 4(1 R_i)\mathbb{1}\{A_{2,i} = a_2\})$

•
$$\mathbf{D}^{(d)}(\mathbf{X}_i) = rac{\partial}{\partial(\eta^{ op}, \gamma^{ op})^{ op}} \mu^{(d)}(\mathbf{X}_i; \eta, \gamma)$$

 $m{\cdot}$ $m{V}^{(d)}\left(m{ au}
ight)$ is a working model for $m{V}$ ar $\left(m{Y}^{(d)}-m{\mu}^{(d)}(m{X}_{i};m{\eta},m{\gamma})\midm{X}_{i}
ight)$

. Lu, X., et al. (2016). Stat. Med.

Goal:

For this analysis, develop a sample size formula for SMARTs with a continuous longitudinal outcome in which the primary aim is to compare, at end-of-study, two embedded DTRs which recommend different first-stage treatments.

Context:

- Three timepoints
- Randomization probability 0.5
- Exchangeable correlation structure

$$N \geq \frac{4\left(Z_{1-\alpha/2} + Z_{1-\beta}\right)^2}{\delta^2} \cdot (1-\rho^2) \cdot (2-r)$$

where

•
$$\delta = \mathsf{E}[\mathsf{Y}_2^{(d)} - \mathsf{Y}_2^{(d')}] / \sqrt{\left(\mathsf{Var}(\mathsf{Y}_2^{(d)}) + \mathsf{Var}(\mathsf{Y}_2^{(d')})\right) / 2}$$
 is the

- α is the desired type-I error
- 1 $-\beta$ is the desired power
- $\rho = cor(Y_t, Y_{t'})$ for $t \neq t'$
- $r = P(R_i = 1)$

$$N \geq \underbrace{\frac{4\left(\mathbf{z}_{1-\alpha/2} + \mathbf{z}_{1-\beta}\right)^2}{\delta^2}}_{\text{Standard sample size for a 2-arm trial}} \cdot (1-\rho^2) \cdot (2-r)$$

where

•
$$\delta = \mathsf{E}[Y_2^{(d)} - Y_2^{(d')}] / \sqrt{\left(\mathsf{Var}(Y_2^{(d)}) + \mathsf{Var}(Y_2^{(d')})\right) / 2}$$
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$$N \geq \frac{4\left(Z_{1-\alpha/2} + Z_{1-\beta}\right)^2}{\delta^2} \cdot \underbrace{(1-\rho^2)}_{\text{Deflation for repeated measures}} \cdot (2-r)$$

where

•
$$\delta = \mathsf{E}[\mathsf{Y}_2^{(d)} - \mathsf{Y}_2^{(d')}] / \sqrt{\left(\mathsf{Var}(\mathsf{Y}_2^{(d)}) + \mathsf{Var}(\mathsf{Y}_2^{(d')})\right) / 2}$$
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$$N \geq \frac{4\left(Z_{1-\alpha/2} + Z_{1-\beta}\right)^2}{\delta^2} \cdot (1-\rho^2) \cdot \underbrace{(2-r)}_{\text{Inflation for SMART design}}$$

where

•
$$\delta = \mathsf{E}[Y_2^{(d)} - Y_2^{(d')}] / \sqrt{\left(\mathsf{Var}(Y_2^{(d)}) + \mathsf{Var}(Y_2^{(d')})\right) / 2}$$
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Table 1: Example sample sizes for comparison of two embedded DTRs. r = 0.4, $\alpha = 0.05$ (two-sided), and $1 - \beta = 0.8$.

	W	Within-Person Correlation			
Std. Effect Size	$\rho = 0$	ho= 0.3	ho = 0.6		
$\delta=$ 0.3	559	508	358		
$\delta = 0.5$	201	183	129		

Working Assumptions for Sample Size

1. Response is uncorrelated with products of first-stage residuals. For any $t_i \leq t_i \leq t^*$,

$$\mathsf{Cov}\left(\mathsf{R}^{(a_1)}, \left(\mathsf{Y}^{(d)}_{t_i} - \mu^{(d)}_{t_i}\right)\left(\mathsf{Y}^{(d)}_{t_j} - \mu^{(d)}_{t_j}\right)\right) = \mathsf{O}$$

[.] Oetting, A. I., et al. (2011).

Working Assumptions for Sample Size

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2. Constrained conditional covariances.

2.1
$$E\left[\left(Y_{2}^{(d)}-\mu_{2}^{(d)}\right)^{2} \mid R^{(a_{1})}=0\right] \leq Var\left(Y_{2}^{(d)}\right)$$

2.2 $Cov(Y_{t}^{(d)},Y_{2}^{(d)}\mid R=1) \leq Cov(Y_{t}^{(d)},Y_{2}^{(d)}\mid R=0)$ for all d and $t=0,1$.

. Oetting, A. I., et al. (2011).

3. Exchangeable correlation structure.

$$\operatorname{Var}\left(\mathbf{Y}^{(d)}\right) = \sigma^{2} \begin{bmatrix} 1 & \rho & \rho \\ \rho & 1 & \rho \\ \rho & \rho & 1 \end{bmatrix}$$

for all d.

Target: $1 - \beta$ = 0.8, α = 0.05 (two-sided)

		Empirical power						
δ	P(R = 1)	ρ	Ν	All satisfied	1 violated	2.1 violated	2.2 violated	
0.3	0.4	0	559	0.801	0.778*	0.803	-	
		0.3	508	0.804	0.800	0.797	0.798	
		0.6	358	0.817	0.807	0.759*	0.788	
		0.8	201	0.836	0.809	-	0.792	
	0.6	0	489	0.804	0.736*	0.810	-	
		0.3	445	0.797	0.758*	0.795	0.780*	
		0.6	313	0.824	0.793	0.752*	0.770*	
		0.8	176	0.845	0.754*	-	0.776*	

* Result is significantly less than 0.8 at the 0.05 significance level.

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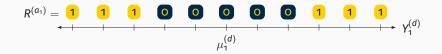


Working Assumptions for Sample Size

1. Response is uncorrelated with products of first-stage residuals. For any $t_i \leq t_j \leq t^*$,

$$\operatorname{Cov}\left(R^{(a_1)},\left(Y_{t_i}^{(d)}-\mu_{t_i}^{(d)}\right)\left(Y_{t_j}^{(d)}-\mu_{t_j}^{(d)}\right)\right)=0$$

Intuition: If this is not true, the relationship between, say $Y_1^{(d)}$ and *R* might look like this:



Two Definitions of Response

$$R^{(a_{1})} = \mathbb{1}\left\{\left(Y_{1}^{(d)}\right)^{2} > 4.7\right\}$$

$$R^{(a_{1})} = \mathbb{1}\left\{Y_{1}^{(d)} > 0.7\right\}$$

$$R^{(a_{1})} = \mathbb{1}\left\{Y_{1}^{(d)} > 0.7\right\}$$

$$R^{(a_{1})} = \mathbb{1}\left\{X_{1}^{(d)} > 0.7\right\}$$

$$R^{(a_{1})} = \mathbb{1}\left\{R^{(a_{1})} = \mathbb{1}\right\}$$

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